

## Nowcasting food inflation with a massive amount of online prices

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*The views expressed herein are those of their authors and not necessarily the views of Narodowy Bank Polski.*

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# Outline

- 1 Introduction
- 2 Literature review
- 3 Methodology
- 4 Results
- 5 Conclusions

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## Summary

- 1 Precise nowcasts are crucial for providing accurate forecasts (Faust and Wright, 2013).
- 2 Analysing online prices gaining a lot of momentum in recent years.
- 3 The research is spearheaded by studies within the Billion Prices Project, launched at MIT in 2008 (Cavallo & Ribogon, 2016).
- 4 We show that online prices can be very effective in nowcasting inflation and macroeconomic practitioners can make use of them already after a couple of months of data collection.

# This paper

- 1 We focus on inflation nowcasting.
- 2 We employ a unique, extensive dataset of online food and non-alcoholic beverages prices gathered from web since 2009.
- 3 Our database contains 159 millions of prices for around 640 thousands of products.
- 4 We perform a real-time nowcasting experiment among popular, simple univariate approaches using highly disaggregated framework.

## Our contribution

### ■ In the paper we:

- 1 discuss how macroeconomic practitioners can improve inflation nowcasts.
- 2 provide evidence for a large number of highly disaggregated inflation components.
- 3 study the usefulness of online prices for a small, emerging economy.
- 4 report how forecasting errors are related to the:
  - scale of the project, i.e. the number of stores being web scraped,
  - the approach towards data curation, i.e. the classification of products,
  - the variability in consumer prices.
- 5 examine the accuracy of nowcasting with online prices during COVID-19 pandemic.

## Key results

### ■ We show that:

- 1 pure estimates of online price changes is already effective in nowcasting food inflation.
- 2 incorporating information on online prices into model-based frameworks delivers a substantial increase in the nowcast accuracy.
- 3 this approach outperforms a variety of frameworks, including judgemental methods.
- 4 marked improvement can be obtained for a number of inflation components, especially those experiencing high volatility throughout the year.
- 5 during COVID-19 the nowcasting quality relatively improved and remained comparable with judgemental nowcasts.
- 6 more data is not always better – product selection and expenditure weighing is essential for providing accurate both in-sample fit as well as out-of-sample nowcasts

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## Usefulness of online data

- Literature on online prices is mushrooming (Lunnemann and Wintr, 2011; Cavallo, 2013; Cavallo & Ribogon, 2016; Cavallo, 2017; Gorodnichenko & Talavera, 2017; Cavallo, 2018; Gorodnichenko et al., 2018).
- Online prices are increasingly included in the compilation of the CPI (in the US, the UK, the Netherlands, New Zealand and Norway,).
- Evidence on the usefulness of scraped data in forecasting inflation remains scarce (Aparicio & Bertolotto, 2020).
  - Scraped data from July 2008 to September 2016.
  - Data for 10 advanced economies.
  - Parsimonious models with online prices beat traditional benchmarks and two leading survey of professional forecasters.
- Our work is conceptually similar but we:
  - provide evidence for a large number of highly disaggregated components,
  - study the usefulness of online prices for an emerging economy with the online market much less developed,
  - show how forecasting errors are related to the scale of the project, the approach towards data curation and the variability in online prices,
  - examine the nowcast accuracy during COVID-19.

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## Online data

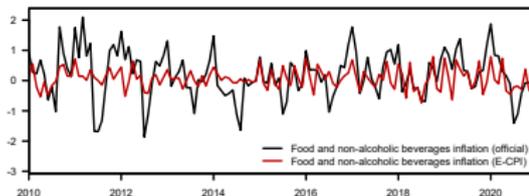
- E-CPI project aimed at collecting online prices launched in December, 2009.
- Frequency of operation changed to daily since 2017.
- Systematically expanding product categories to account for clothing, footwear, products, electronics, drugs, air plain tickets.

Store ID	$q$	$q_s$	$n_p$	$\bar{n}_p$	$\bar{n}_d$	$med\{n_d\}$
1	82,995	19,250	13,394,437	7,327	161	42
2	56,739	16,368	14,721,627	8,544	260	76
3	76,652	13,007	25,600,013	17,534	334	222
4	50,954	15,935	17,196,321	11,495	341	164
5	132,543	26,175	43,387,750	23,709	327	187
6	21,658	7,381	6,346,966	9,544	302	266
7	103,234	34,054	16,708,475	9,170	175	87
8	119,168	27,544	21,630,105	13,213	182	34
Total	643,943	159,714	158,985,694	12,759	250	102

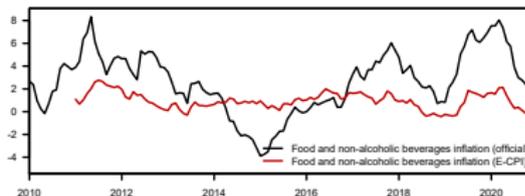
$q_s$  – number of selected products,  $n_p$  – number of observed prices,  $\bar{n}_p$  – average number of prices per day,  $\bar{n}_d$  – mean number of days price is observed,  $med\{n_d\}$  – median number of days price is observed.

# In-sample tracking accuracy of online data

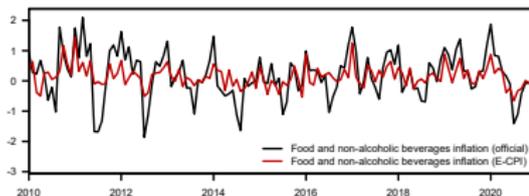
**A. No product selection (monthly dynamics)**



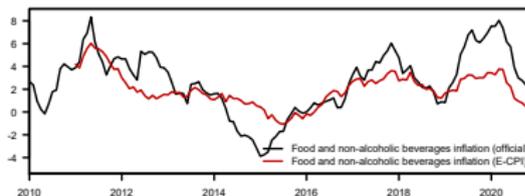
**B. No product selection (yearly dynamics)**



**C. Simple selection and official aggregation (monthly dynamics)**



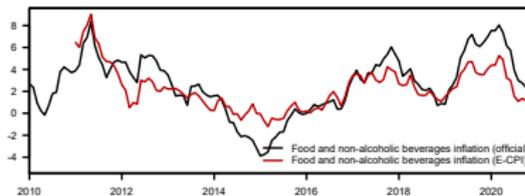
**D. Simple selection and official aggregation (yearly dynamics)**



**E. Official product selection and aggregation (monthly dynamics)**



**F. Official product selection and aggregation (monthly dynamics)**



## Nowcasting competition (1)

- Inflation rate as the monthly, non-seasonally adjusted change in prices as the variable of interest.
- Change in the online price index defined as  $o_{c,t}$ .
- Baseline approach – recursive estimation strategy making use of expanding window.
  - Estimation sample spanning the period January 1999 - December 2016.
  - Evaluation sample spanning the period January 2017 - December 2020.
- Nowcast evaluation with MFE, RMSFE and Diebold-Mariano tests.
  - MFE reported in levels.
  - RMSFE reported as ratios - a value above 1 (below 1) that the competing approach produces on average less (more) accurate nowcasts than  $EC^{SX}$ .
- Sensitivity analysis for:
  - rolling window estimation,
  - store composition,
  - data curation,
  - forecast combinations,
  - the COVID-19 period (i.e. January 2020 - December 2020).

## Nowcasting competition (2)

### ■ Model entering the nowcasting competition:

- 1 The simple random walk:  $p_{c,t+1} = p_{c,t}$ .
- 2 The random walk à la Atkeson & Ohanian (2001):  $p_{c,t+1} = \frac{1}{12} \sum_{j=1}^{12} p_{c,t-j+1}^{SA} + \hat{s}_{c,t+1}^{TS}$ .
- 3 The best SARMA model based on in-sample fit ( $BS^{IS}$ ):  $p_{c,t} = \mu_c + \varepsilon_{c,t} + \sum_{i=1}^P \phi_{c,i} p_{c,t-i} + \sum_{i=1}^Q \theta_{c,i} \varepsilon_{c,t-i} + \sum_{i=1}^P \Phi_{c,i} p_{c,t-i*12} + \sum_{i=1}^Q \Theta_{c,i} \varepsilon_{c,t-i*12}$ , assuming that  $\varepsilon_{c,t} \sim NIID(0, \sigma_c^2)$  chosen from 64 specifications using BIC.
- 4 The best SARMA model based on out-of-sample accuracy ( $BS^{OS}$ ) using the minimal RMSFE criterion calculated on a pseudo validation set.
- 5 The pure, ex-post E-CPI nowcast:  $p_{c,t+1} = o_{c,t+1}$ .
- 6 The pure, real-time E-CPI nowcast ( $EC^{RT}$ ):  $p_{c,t+1} = o_{c,t+1}^*$ .
- 7 The best SARMAX model with online prices as the exogenous variable ( $EC^{SX}$ ). The specification is the following:  $p_{c,t} = \mu_c + \varepsilon_{c,t} + \sum_{i=0}^B \beta_{c,i} o_{c,t-i} + \sum_{i=1}^P \phi_{c,i} p_{c,t-i} + \sum_{i=1}^Q \theta_{c,i} \varepsilon_{c,t-i} + \sum_{i=1}^P \Phi_{c,i} p_{c,t-i*12} + \sum_{i=1}^Q \Theta_{c,i} \varepsilon_{c,t-i*12}$  and  $\varepsilon_{c,t} \sim NIID(0, \sigma_c^2)$  chosen from 256 specifications using the minimal RMSFE criterion calculated on a pseudo validation set.
- 8 The simple combination of the  $EC^{SX}$  models ( $EC^{MC}$ ).
- 9 The combination of  $BS^{OS}$  models with equal weights ( $BS^{MC}$ ).
- 10 The combination of  $EC^{SX}$  models with weight inversely proportional to RMSFE ( $EC^{MCI}$ ).
- 11 The judgemental forecast ( $JD$ ).

# Outline

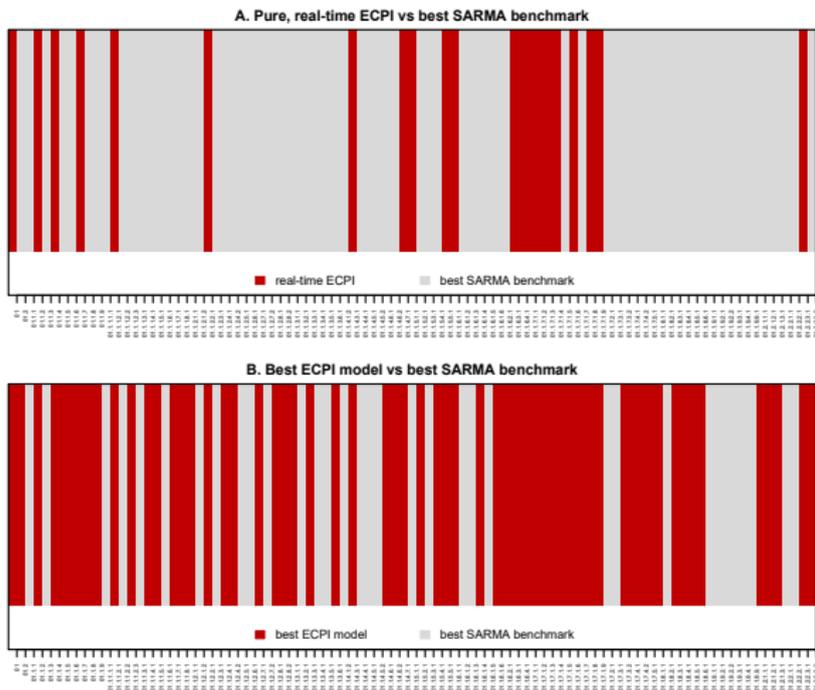
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- 3 Methodology
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## Baseline results – the aggregate

Error	Online prices			Traditional benchmarks				Forecast combinations		
	$EC^{SX}$	$EC^{EX}$	$EC^{RT}$	$RW$	$AO^{SA}$	$BS^{IS}$	$BS^{OS}$	$BS^{MC}$	$EC^{MC}$	$EC^{MCI}$
Recursive estimation										
MFE	0.018	-0.106	-0.093	0.042	0.118	0.109	0.111	0.134	0.025	0.022
RMSFE	0.338	1.071	1.082	1.784 <sup>a</sup>	1.528 <sup>a</sup>	1.446 <sup>a</sup>	1.524 <sup>a</sup>	1.621 <sup>a</sup>	1.181 <sup>b</sup>	1.176 <sup>b</sup>
Rolling estimation										
MFE	0.019	-0.106	-0.093	0.042	0.118	0.041	0.017	0.025	0.080	0.019
RMSFE	0.427	0.848	0.857	1.413 <sup>b</sup>	1.209 <sup>c</sup>	1.362 <sup>b</sup>	1.381 <sup>b</sup>	1.324 <sup>b</sup>	1.003	0.965
COVID-19										
MFE	0.146	-0.076	-0.078	0.084	0.362	0.401	0.394	0.394	0.212	0.214
RMSFE	0.340	1.312	1.332	1.567	1.575	1.727	1.902	2.048	1.411	1.398

MFE reported in levels. RMSFE reported as ratios - a value above 1 (below 1) that the competing approach produces on average less (more) accurate nowcasts than  $EC^{SX}$ . <sup>a</sup> denotes significance at the 1 percent level, <sup>b</sup> denotes significance at the 5 percent level, <sup>c</sup> denotes significance at the 10 percent level.

## Baseline results – low level disaggregation

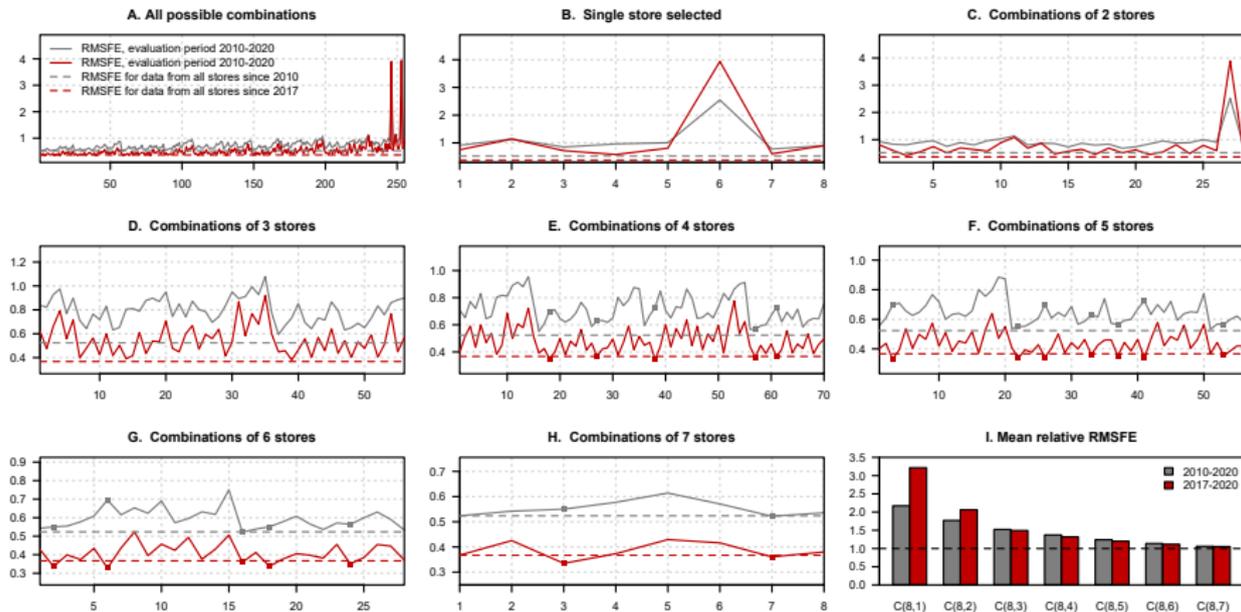


## The comparison between online data approach and judgemental nowcasts

	Recursive estimation whole evaluation period		Rolling estimation		Recursive estimation COVID-19 period	
	$EC^{SX}$	$JD$	$EC^{SX}$	$JD$	$EC^{SX}$	$JD$
MFE	0.018	0.021	0.019	0.021	0.146	0.147
RMSFE	0.338	1.144 <sup>c</sup>	0.427	0.905	0.340	0.965

MFE reported in levels. RMSFE reported as ratios - a value above 1 (below 1) that the competing approach produces on average less (more) accurate nowcasts than  $EC^{SX}$ . <sup>a</sup> denotes significance at the 1 percent level, <sup>b</sup> denotes significance at the 5 percent level, <sup>c</sup> denotes significance at the 10 percent level.

# Is more data better?



## Not always – the importance of data curation

		RMSFE	MFE
Frameworks with online prices	$EC^{SX}$	0.338	0.018
	$EC^{EX}$	1.071	-0.106
	$EC^{RT}$	1.082	-0.093
Simple selection	$EC^{SX}$	1.299 <sup>b</sup>	0.036
	$EC^{RT}$	1.400 <sup>a</sup>	-0.117
Traditional benchmarks	$RW$	1.784 <sup>a</sup>	0.042
	$AO^{SA}$	1.528 <sup>a</sup>	0.118
	$BS^{IS}$	1.446 <sup>a</sup>	0.109
	$BS^{OS}$	1.524 <sup>a</sup>	0.111
Forecast combinations	$BS^{MC}$	1.621 <sup>a</sup>	0.134
	$EC^{MC}$	1.181 <sup>b</sup>	0.025
	$EC^{MCI}$	1.176 <sup>b</sup>	0.022

MFE reported in levels. RMSFE reported as ratios - a value above 1 (below 1) that the competing approach produces on average less (more) accurate nowcasts than  $EC^{SX}$ . Simple selection refers to the robustness check, where products are classified into respective groups using unsupervised learning based on word stems. <sup>a</sup> denotes significance at the 1 percent level, <sup>b</sup> denotes significance at the 5 percent level, <sup>c</sup> denotes significance at the 10 percent level.

## The impact of price dispersion on the relative nowcast accuracy

	$RMSFE^{OS \rightarrow SX^*}$	$RMSFE^{SX^* \rightarrow SX}$	$RMSFE^{OS \rightarrow SX}$
log(SD)	-0.014 (0.011)	-0.046*** (0.012)	-0.059*** (0.013)
C01.1.1	0.042 (0.035)	0.064* (0.035)	0.099** (0.046)
C01.1.2	0.079** (0.031)	0.052 (0.036)	0.122*** (0.041)
C01.1.3	0.088*** (0.027)	0.052 (0.033)	0.134*** (0.040)
C01.1.4	0.053 (0.041)	0.056 (0.036)	0.102** (0.043)
C01.1.5	0.033 (0.033)	0.116*** (0.042)	0.139*** (0.045)
C01.1.6	0.089** (0.044)	0.109* (0.056)	0.188*** (0.070)
C01.1.7	0.039 (0.034)	0.074 (0.045)	0.107** (0.051)
C01.1.8	0.053** (0.025)	0.057 (0.040)	0.104*** (0.044)
C01.1.9	0.097*** (0.023)	0.053 (0.033)	0.144*** (0.037)
C01.2	0.046 (0.033)	0.033 (0.037)	0.072* (0.043)
Constant	-0.082*** (0.021)	-0.079** (0.033)	-0.154*** (0.038)
Observations	95	95	95
R <sup>2</sup>	0.156	0.311	0.401
F Statistic (df = 11; 83)	1.391	3.412***	5.055***

The table reports the outcomes of a simple regression of the dispersion in consumer prices measured with the standard deviation across all inflation components on the relative RMSFE measures, after accounting for group-specific fixed effects. C01.1.1 denotes *Bread and cereals*, C01.1.2 denotes *Meat*, C01.1.3 denotes *Fish and seafood*, C01.1.4 denotes *Milk, cheese and eggs*, C01.1.5 denotes *Oils and fats*, C01.1.6 denotes *Fruits*, C01.1.7 denotes *Vegetables*, C01.1.8 denotes *Sugar, jam, honey, chocolate and confectionery*, C01.1.9 denotes *Food products, n.e.c.* and C01.2 denotes *Non-alcoholic beverages*. Robust standard errors are reported in parentheses. \*\*\* denotes significance at the 1 percent level, \*\* denotes significance at the 5 percent level, \* denotes significance at the 10 percent level.

## Recursive versus rolling estimation

		$EC^{SX}$	$EC^{MC}$	$EC^{MCI}$	$BS^{IS}$	$BS^{OS}$	$BS^{MC}$	$BS^{MCI}$
RMSFE	rolling window	0.427	0.428	0.412	0.581	0.589	0.565	0.566
	expanding window	0.338	0.399	0.397	0.489	0.515	0.548	0.545
	ratio	26%	7%	4%	19%	14%	3%	4%
MFE	rolling	0.019	0.080	0.019	0.041	0.017	0.025	0.027
	expanding	0.018	0.025	0.022	0.109	0.111	0.134	0.131
	difference	0.001	0.054	-0.003	-0.068	-0.093	-0.109	-0.104

For the RMSFE statistics the ratio denotes the percent change in the RMSFE when the expanding windows estimation is switched to the rolling windows. For the MFE a simple difference is reported.

## Nowcasting during the COVID-19 period (2020M1-2020M12)

Error	Online prices			Traditional benchmarks				Forecast combinations		
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## Main takeaways

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- 2 Incorporating information on online prices into model-based frameworks delivers a substantial increase in the nowcast accuracy.
- 3 This approach outperforms a variety of frameworks, including judgemental methods.
- 4 Marked improvement can be obtained for a number of inflation components, especially those experiencing high volatility throughout the year.
- 5 During COVID-19 the nowcasting quality relatively improved and remained comparable with judgemental nowcasts.
- 6 Meticulous product selection and expenditure weighing is essential for providing accurate both in-sample fit as well as out-of-sample nowcasts.